

Le Cam, Birgé, Huber, and Strassen

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Total Variation and Hellinger Distances



My first two years of studies revolve around the total variation and Hellinger distances.

$$\text{TV}(p, q) := \frac{1}{2} \int |p - q|,$$

$$H(p, q) := \sqrt{\frac{1}{2} \int (\sqrt{p} - \sqrt{q})^2}.$$

- First year: $\text{TV}(f_\pi, f_\eta) \stackrel{?}{\asymp} H(f_\pi, f_\eta)$ for Gaussian location mixtures.

$$f_\pi(x) := \int_{\mathbb{R}^d} \phi_d(x - \theta) d\pi(\theta).$$

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- Second year (ongoing): Robust estimation of a location parameter (under Huber's contamination model).

First Year: Resolving an Open Question



- An open question [JPW23]: $TV(f_\pi, f_\eta) \stackrel{?}{\asymp} H(f_\pi, f_\eta)$ for $\text{supp}(\pi), \text{supp}(\eta) \subseteq [-M, M]^d$.

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- Our answer [JG26]: No. We proved that

$$H(f_\pi, f_\eta) \leq \text{TV}(f_\pi, f_\eta)^{1 - \frac{\Theta(1)}{\log \log(1/\text{TV}(f_\pi, f_\eta))}}.$$

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$$H(f_\pi, f_\eta) \leq \text{TV}(f_\pi, f_\eta)^{1 - \frac{\Theta(1)}{\log \log(1/\text{TV}(f_\pi, f_\eta))}}.$$

- As a consequence, for $X_1, \dots, X_n \stackrel{i.i.d.}{\sim} (1 - \epsilon)P_{f_\pi} + \epsilon Q$,

$$\inf_{\hat{f}} \sup_{\pi, Q} \mathbb{E}_{\epsilon, \pi, Q} \left[H^2 \left(f_\pi, \hat{f} \right) \right] \asymp \epsilon^{2 \left(1 - \frac{\Theta(1)}{\log \log(1/\epsilon)} \right)},$$

in the “large- ϵ regime” where $\epsilon \gtrsim n^{-0.49}$.

Motivation: Gaussian Location Model (1/2)



Recall the simplest model in robust statistics.

- Classical result: For $X_1, \dots, X_n \stackrel{i.i.d.}{\sim} (1 - \epsilon)N(\theta, 1) + \epsilon Q$,

$$\inf_{\hat{\theta}} \sup_{\theta, Q} \inf \left\{ \rho > 0 : \mathbb{P}_{\epsilon, \theta, Q} \left[(\hat{\theta} - \theta)^2 \geq \rho \right] \leq 0.05 \right\} \asymp \frac{1}{n} + \epsilon^2,$$

uniformly for all n and $\epsilon \in [0, 0.49]$.

- The Gaussian location model under Huber's contamination admits two regimes:

$\frac{1}{n}$ ← Small- ϵ regime

ϵ^2 ← Large- ϵ regime



Motivation: Gaussian Location Model (2/2)

What determines the minimax rates in the two regimes?

$$\left(\widehat{\theta} - \theta\right)^2 \asymp \frac{1}{n} + \epsilon^2,$$

- The first term is characterized by the Hellinger distance.

$$r^2(n) := \sup \left\{ \theta^2 : H^2(N(0, 1), N(\theta, 1)) \leq \frac{1}{n} \right\} \asymp \frac{1}{n}.$$



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- The second term is given by the total variation.

$$\begin{aligned} \omega^2(\epsilon) &:= \sup \left\{ \theta^2 : \exists Q_0, Q_1, \right. \\ &\quad \left. (1 - \epsilon)N(0, 1) + \epsilon Q_0 = (1 - \epsilon)N(\theta, 1) + \epsilon Q_1 \right\} \\ &= \sup \left\{ \theta^2 : \text{TV}(N(0, 1), N(\theta, 1)) \leq \frac{\epsilon}{1 - \epsilon} \right\} \asymp \epsilon^2. \end{aligned}$$

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- We cannot distinguish θ from zero even when we observe $n = \infty$ sample points.

Second Year: Singular Location Model (1/4)



What happens under singular location models?

- Infinite Fisher information
- Example: Density plot of $\Delta(0,1)$



Second Year: Singular Location Model (2/4)



First follow the same argument as the Gaussian location model.

- For $X_1, \dots, X_n \stackrel{i.i.d.}{\sim} (1 - \epsilon)\Delta(\theta, 1) + \epsilon Q$,

$$\left(\widehat{\theta} - \theta\right)^2 \stackrel{?}{\asymp} r^2(n) + \omega^2(\epsilon),$$

uniformly for all n and $\epsilon \in [0, 0.49]$, where

$$r^2(n) := \sup \left\{ \theta^2 : H^2(\Delta(0, 1), \Delta(\theta, 1)) \leq \frac{1}{n} \right\},$$

$$\omega^2(\epsilon) := \sup \left\{ \theta^2 : \text{TV}(\Delta(0, 1), \Delta(\theta, 1)) \leq \frac{\epsilon}{1 - \epsilon} \right\}.$$

Second Year: Singular Location Model (3/4)



- The first term is characterized by the Hellinger distance.

$$r^2(n) := \sup \left\{ \theta^2 : \underbrace{H^2(\Delta(0, 1), \Delta(\theta, 1))}_{\asymp \theta^2 \log \frac{1}{\theta}} \leq \frac{1}{n} \right\} \asymp \frac{1}{n \log n}.$$

- The second term is given by the total variation.

$$\begin{aligned} \omega^2(\epsilon) &:= \sup \left\{ \theta^2 : \text{TV}(\Delta(0, 1), \Delta(\theta, 1)) \leq \frac{\epsilon}{1 - \epsilon} \right\} \\ &= 4 \left(1 - \sqrt{\frac{1 - 2\epsilon}{1 - \epsilon}} \right)^2 \asymp \epsilon^2. \end{aligned}$$

Second Year: Singular Location Model (4/4)



- For $X_1, \dots, X_n \stackrel{i.i.d.}{\sim} (1 - \epsilon)\Delta(\theta, 1) + \epsilon Q$,

$$\left(\hat{\theta} - \theta\right)^2 \stackrel{?}{\asymp} \frac{1}{n \log n} + \epsilon^2,$$

Second Year: Singular Location Model (4/4)



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- Our answer: No. We proved that

$$\left(\hat{\theta} - \theta\right)^2 \asymp \frac{1}{n \log n} + \underbrace{\frac{1}{n \log \left(\frac{1}{n\epsilon^2} \vee e\right)}}_{\text{Intermediate-}\epsilon} + \epsilon^2,$$

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- What is going on?

The Intermediate- ϵ Regime (1/3)



We recall the classical results for composite-composite hypothesis testing: \mathfrak{H}_0^ϵ v.s. $\mathfrak{H}_\theta^\epsilon$, where we define $\mathfrak{H}_\theta^\epsilon := \{(1 - \epsilon)F_\theta + \epsilon Q : Q\}$.

- Le Cam, Lucien & Birgé, Lucien

$$\begin{aligned} 1 - \sqrt{2nH^2(\mathfrak{H}_0^\epsilon, \mathfrak{H}_\theta^\epsilon)} &\leq \inf_{\phi} \sup_{(P_0, P_1) \in \mathfrak{H}_0^\epsilon \times \mathfrak{H}_\theta^\epsilon} P_0^{\otimes n} \phi + P_1^{\otimes n} (1 - \phi) \\ &\leq 2 \exp(-nH^2(\mathfrak{H}_0^\epsilon, \mathfrak{H}_\theta^\epsilon)), \\ H^2(\mathfrak{H}_0^\epsilon, \mathfrak{H}_\theta^\epsilon) &:= \inf_{(P_0, P_1) \in \mathfrak{H}_0^\epsilon \times \mathfrak{H}_\theta^\epsilon} H^2(P_0, P_1). \end{aligned}$$

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- Huber, Peter & Strassen, Volker

$$\begin{aligned} \inf_{\phi} \sup_{(P_0, P_1) \in \mathfrak{H}_0^\epsilon \times \mathfrak{H}_\theta^\epsilon} P_0^{\otimes n} \phi + P_1^{\otimes n} (1 - \phi) &= 1 - \text{TV}((P_0^*)^{\otimes n}, (P_1^*)^{\otimes n}), \\ H^2(P_0^*, P_1^*) &= H^2(\mathfrak{H}_0^\epsilon, \mathfrak{H}_\theta^\epsilon), \\ dP_1^*/dP_0^* &= \max(c', \min(c'', dF_\theta/dF_0)). \end{aligned}$$

The Intermediate- ϵ Regime (2/3)



- According to Le Cam, Birgé, Huber, and Strassen,

$$\delta^2(n, \epsilon) := \sup \left\{ \theta^2 : H^2(\mathfrak{H}_0^\epsilon, \mathfrak{H}_\theta^\epsilon) \leq \frac{1}{n} \right\},$$
$$\mathfrak{H}_\theta^\epsilon := \{(1 - \epsilon)\Delta(\theta, 1) + \epsilon Q : Q\}.$$

- We proved that

$$H^2(\mathfrak{H}_0^\epsilon, \mathfrak{H}_\theta^\epsilon) \asymp \begin{cases} \theta^2 \log \left(\frac{\theta}{\theta^2 \vee \epsilon} \vee e \right), & \omega(\epsilon) < \theta, \\ 0, & \omega(\epsilon) \geq \theta. \end{cases}$$

- Thus, solving $nH^2 \asymp 1$ gives

$$\delta^2(n, \epsilon) \asymp \frac{1}{n \log n} + \frac{1}{n \log \left(\frac{1}{n\epsilon^2} \vee e \right)} + \epsilon^2.$$

The Intermediate- ϵ Regime (3/3)

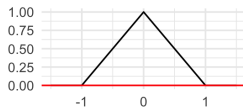


Density

Shape

$\delta^2(n, \epsilon)$

Triangle



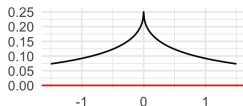
$$\frac{1}{n \log n} + \frac{1}{n \log\left(\frac{1}{nc^2} \vee e\right)} + \epsilon^2$$

Semicircle



$$\frac{1}{n^{4/3}} + \left(\frac{\epsilon}{n}\right)^{2/3} + \epsilon^2$$

GG(1/2)



$$\frac{1}{n \log n} + \frac{1}{n \log\left(\frac{1}{nc^2} \vee e\right)} + \epsilon^2$$

Conclusion



- Such intermediate- ϵ regimes exist only under **singular** models.
- Q: What if we have more than one singularities?
- Q: Unified framework?
 - Le Cam, Lucien & Birgé, Lucien

$$nH^2(P_0^*, P_1^*) \asymp 1.$$

- Huber, Peter & Strassen, Volker

$$dP_1^*/dP_0^* = \max(c', \min(c'', dF_\theta/dF_0)).$$

Acknowledgements



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- I thank my wife.



-  Joonhyuk Jung and Chao Gao, *Sharp inequalities between total variation and Hellinger distances for Gaussian mixtures*, arXiv preprint arXiv:2602.03202 (2026).
-  Zeyu Jia, Yury Polyanskiy, and Yihong Wu, *Entropic characterization of optimal rates for learning Gaussian mixtures*, The Thirty Sixth Annual Conference on Learning Theory, PMLR, 2023, pp. 4296–4335.